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Section: 8-9

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Question 1 (5 pts)

Given the function:  $G_{XY}(x,y) = u(x)u(y)(1-\exp(-(x+y)))$ .

Is this a density or distribution function? distribution

Is this a valid probability function? Yes

$1 - e^{-(x+y)} = 1 - (e^{-x}e^{-y})$

Why? because it satisfies the conditions (1)

For  $x > 0$  or  $y > 0$

$G_{XY}(x,y) = 1 - e^{-(x+y)} \leq 1$  because  $1 - e^{-(x+y)} \rightarrow 1 - 1 = 0$  min

condition (1)

$G_{XY}(0,y) = 1 - e^{-y} = 1$

$G_{XY}(x,\infty) = 1$

$G_{XY}(\infty,0) = 1$

condition (2)

$P(x_a < X \leq x_b, y_a < Y \leq y_b) = F_{X,Y}(x_b, y_b) + F_{X,Y}(x_a, y_a) - F_{X,Y}(x_b, y_a) - F_{X,Y}(x_a, y_b)$

is between 0 and 1 from figure and it's not density because area under it is not 1

Question 2 (5 pts)

Find the constant b (in terms of a) for the function:  $f_{XY}(x,y) = b \exp(-(x+y))$  for  $0 < x < a$  and  $0 < y < \infty$ , and zero elsewhere, so that it is a valid density function

$\int_0^a \int_0^{\infty} b e^{-(x+y)} dx dy = 1 \rightarrow -b \left[ \frac{e^{-(a+y)}}{-1} - \frac{e^{-y}}{-1} \right]_0^a = -b \left[ \frac{e^{-(a+y)}}{-1} - \frac{e^{-y}}{-1} \right]_0^a$

$b = \frac{1}{1 - e^{-a}}$

$= -b \left[ -e^{-(a+y)} + e^{-y} \right]_0^a$   
 $= -b [ 0 + 0 - (-e^{-a} + 1) ]$   
 $= b(1 - e^{-a})$  now  $b(1 - e^{-a}) = 1$

Find the distribution function

$f_{X,Y}(x,y) = \begin{cases} 0 & x < 0 \text{ or } y < 0 \\ b e^{-(x+y)} & 0 < x < a \text{ and } 0 < y < \infty \\ b e^{-x} + b & a < x \text{ and } 0 < y < \infty \end{cases}$

$F_{XY}(x,y)$

$b = \frac{1}{1 - e^{-a}}$

$\int_0^y \int_0^x b e^{-(x+y)} dx dy$

$= -b \left[ \frac{e^{-(x+y)}}{-1} \right]_0^x = -b e^{-(x+y)} \Big|_0^x = -b e^{-x-y} + b e^{-y}$

$= \int_0^y (-b e^{-x-y} + b e^{-y}) dy$

$= \left( +b e^{-x-y} - b e^{-y} \right) \Big|_0^y$

$= b e^{-x-y} - b e^{-y} - (b e^{-x} - b)$

$$\int_{y_a}^{y_b} \int_{-\infty}^{\infty} f_{X,Y}$$

$$\int_{y_a}^{y_b} f_{Y|X}$$

$$x e^{-x(1+y)} - \frac{-x(1+y)}{(1+y)^2}$$

$$= 0 - 0 - (0 - \frac{1}{(1+y)^2}) = \frac{1}{(1+y)^2}$$

$$x \frac{-x(1+y)}{e^{-x(1+y)}} = \frac{-x^2(1+y)}{e^{-x(1+y)}}$$

$$= \frac{-x^2(1+y)}{-x(1+y)} = -x$$

Question 3, (5 pts)  
Find the marginal density functions of the joint function:  
 $f_{X,Y}(x,y) = u(x)u(y) x \exp\{-x(1+y)\}$

$$f_X(x) = \int_{-\infty}^{\infty} u(x)u(y)x e^{-x(1+y)} dy = \int_{-\infty}^{\infty} u(x)x e^{-x(1+y)} dy = x u(x) \frac{e^{-x(1+y)}}{-x} \Big|_{-\infty}^{\infty}$$

$$= -u(x) [0 - e^{-x}] = u(x)e^{-x}$$

$$f_Y(y) = \int_{-\infty}^{\infty} u(x)u(y)x e^{-x(1+y)} dx = \int_{-\infty}^{\infty} u(y)x e^{-x(1+y)} dx = \frac{u(y)}{(1+y)^2}$$

4

$$f_X(x) = u(x)e^{-x}$$

$$f_Y(y) = \frac{u(y)}{(1+y)^2}$$

Find  $P(Y \leq 2 | X=1) = F_Y(2 | X=1) = \int_{-\infty}^2 f_{Y|X}(y | X=1) dy = \int_{-\infty}^2 f_{X,Y}(1,y) dy$

$$\int_{-\infty}^2 u(y) e^{-(1+y)} dy = \int_0^2 e^{-(1+y)} dy = \left[ -e^{-(1+y)} \right]_0^2 = \frac{-e^{-3} - (-e^{-1})}{-e^{-1} + e^{-1}} f_X(1)$$

$$f_X(1) = u(1)e^{-1} = e^{-1}$$

now  $P(Y \leq 2 | X=1) = \frac{e^{-1} - e^{-3}}{e^{-1} - e^{-1}} = 1 - e^{-2}$

Question 4 (5 pts)  
Given  $f_{X,Y}(x,y)$  is a valid joint density function. Find the area under the density function  $f_{Y|X}$

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

$$\text{Area} = \int_{-\infty}^{\infty} f_{Y|X}(y|x) dy = \frac{\int_{-\infty}^{\infty} f_{X,Y}(x,y) dy}{f_X(x)} = \frac{f_X(x)}{f_X(x)} = 1$$

5

Area = 1

$$\frac{1200}{900} = \frac{4}{3}$$

Question 5 (6 pts)

A pointer is spun on a fair wheel of chance numbered from 0 to 60 around its circumference

What is the average value of the pointer position

$$f_x(x) = \begin{cases} \frac{1}{60} & 0 \leq x < 60 \\ 0 & \text{elsewhere} \end{cases}$$

$$\bar{x} = \int_0^{60} x \cdot \frac{1}{60} dx = \frac{x^2}{2(60)} \Big|_0^{60} = \frac{(60)(60)}{2(60)} = 30$$

Average position = 30

What deviation from the mean value will the pointer take on average (RMS deviation from average)

$$m_2 = \int_0^{60} x^2 \cdot \frac{1}{60} dx = \frac{x^3}{3(60)} \Big|_0^{60} = \frac{(60)(60)(60)}{3(60)} = 1200$$

$$\sigma_x^2 = m_2 - m_1^2 = 1200 - (30)^2 = 1200 - 900 = 300$$

$$\sqrt{\sigma_x^2} = \sqrt{300}$$

RMS deviation =  $\sqrt{300}$

Question 6 (5 pts)

A university conducts an entrance exam for the first year students (few thousands). Known is that the exam has 50 questions of different topics (assume the questions are not related). Find the characteristic function and density distribution of the normalized (mean and variance) sum of the total grades. State clearly your evidence.

$$f_x = \sqrt{\frac{2}{\pi}} e^{-x^2/2}$$

~~$$f_y = \frac{1}{\sqrt{2\pi}} e^{-y^2/2}$$~~

$X_1$ : grade of first student  
 $X_2$ : " " " second "  
 $X_3$ : " " " third "  
 $\vdots$   
 $Y$ : sum of total grades

~~$Y$  has a density function~~

$$Y = X_1 + X_2 + X_3 + \dots + X_N$$

$N$  is very large can be described as  $N \rightarrow \infty$  and every  $x_i$  and  $Y$  is not decreasing and we can describe every  $x_i$  that they are independent of each other.

$\therefore Y$  has a density function that is similar to the Gaussian density function.

$$f_y(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2}$$

$$e^{-0} = e^0 = 1$$

$$f_z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$$

$$\text{Phi}_z(w) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} e^{jwy} dy$$

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} e^{jwy} dy$$

Question 7 (5 pts)

Find the 10<sup>th</sup> moment of a Gaussian function of zero mean and unit variance (normalized),  $\sim N(0,1)$

$$m_{10} = \int_{-\infty}^{\infty} x^{10} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx !!$$

(Long)

$$\begin{aligned}
 & x^{10} \\
 & 10 x^9 \\
 & 90 x^8 \\
 & 90(8) x^7 \\
 & 90(8)(7) x^6 \\
 & 90(8)(7)(6)(5) x^5 \\
 & \dots
 \end{aligned}$$

$m_{10} = 0$  Zero

$$\begin{aligned}
 & (-2j\omega)^{10} \\
 & = -1j^{10} \omega^{10} \\
 & = \omega^{10}
 \end{aligned}$$

$$\begin{aligned}
 G_X(\omega) &= \int_{-\infty}^{\infty} e^{j\omega x} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \\
 &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2} + j\omega x} dx \\
 &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \sum_{n=0}^{\infty} \frac{(-\frac{x^2}{2} + j\omega x)^n}{n!} dx \\
 &= \frac{1}{\sqrt{2\pi}} \sum_{n=0}^{\infty} \frac{(-j\omega)^n}{n!} \frac{d^n G_X(\omega)}{d\omega^n} \bigg|_{\omega=0}
 \end{aligned}$$

$x^2$  is dominant against  $x$

$$f_X(x) = \frac{x}{s^2} e^{-\frac{x^2}{2s^2}} \quad x \geq 0$$

non-monotonic

Question 8 (5 pts)

The envelope of a radio detection signal has Rayleigh density.  $f_X(x) = (x/s^2) \exp(-x^2/(2s^2))$ ,  $x \geq 0$ , and you are asked to find the power as  $Y = cX^2$ . Find the density

$$f_Y(y) = \sum_{\text{of } Y} f_X(x_n) \quad \text{where } Y = cX^2 \rightarrow \frac{dY}{dX} = 2cX$$

$$x = \sqrt{\frac{y}{c}} \quad \left| \frac{dx}{dy} \right| = \frac{1}{2\sqrt{cy}}$$

$$f_Y(y) = \frac{f_X(\sqrt{\frac{y}{c}})}{|2c\sqrt{\frac{y}{c}}|} + \frac{f_X(-\sqrt{\frac{y}{c}})}{|-2c\sqrt{\frac{y}{c}}|} = \frac{f_X(\sqrt{\frac{y}{c}}) + f_X(-\sqrt{\frac{y}{c}})}{|2c\sqrt{\frac{y}{c}}|}$$

$$= \frac{2\sqrt{\frac{y}{c}} e^{-y/2cs^2}}{s^2 |2c\sqrt{\frac{y}{c}}|}$$

$$f_Y(y) = \frac{2\sqrt{\frac{y}{c}} e^{-y/2cs^2}}{|2c\sqrt{\frac{y}{c}}|}$$

No aid given, received, nor observed

I swear to Allah

I state, with my word of honor, that I didn't use, or see, any form of cheating, including electronic forms

حرام العلم بالشرف